# STOCHASTIC DIFFERENTIAL EQUATION: AN APPLICATION TO MORTALITY DATA 



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#### Abstract

In the present paper we consider an application of stochastic differential equation to model age-specific mortalities. We use New Zealand mortality data for the period 1948-2015 to fit the model. The point predictions of mortality rates at ages 40,60 and 80 are quite good, almost undistinguishable from the true mortality rates observed.


## 1. INTRODUCTION

Human mortality rates are very important for prediction purposes in social security systems, life insurance, public and private pension plans, etc. The mortality rates of all age and sex groups have a tendency to decline over time, but they clearly show random fluctuations. Many authors have contributed towards the modelling of mortality data through stochastic differential equations (Braumann, 1993; Braumann, 1999a; Braumann, 1999b; Kessler et al., 2012 and Aït-Sahalia, (2002, 2008)). These authors also have contributed towards estimation and other statistical inferential issues on stochastic differential equations. Braumann (1993) and Braumann (1999a) use Black-Scholes model for estimation, testing and prediction, including comparison tests among average return/growth rates of different stocks or populations. Ordinary differential equation (ODE) models have been used to study the growth of individual living beings, like farm animals, trees or fish. When the growth occurs in environments with random variations, stochastic differential equation (SDE) models have been proposed. Braumann (1999b) shows that qualitative results very similar to those obtained for specific models also hold for the general stochastic population growth model.

Lagarto and Braumann (2014) used SDE model (stochastic Malthusian) for mortality data of Portugal and considered precisely the joint evolution of the crude death rates of two age-sex groups. Lagarto (2014) has
considered a study of alternative structures and has determined, among those, which ones have a good performance. These can then be used in predictions and in applications.

## 2. STOCHASTIC MORTALITY MODEL

The random fluctuations are mostly explained by environmental stochasticity and use a stochastic differential equation (SDE) model (Black-Scholes or stochastic Malthusian model) to represent $D(t)$, the death rate at time $t$ of a given age-sex group. In a random environment, the rationale is that the effect of environmental random fluctuations on the growth rate can be described approximately by a white noise, so that its accumulated effect up to time $t$ can be approximated by a Weiner process $\sigma W(t)$. This gives to $\Delta D(t)=(R \Delta t+\sigma \Delta W(t)) D(t)$, with $\Delta W(t)=$ $W(t+\Delta t)-W(t)$, and, letting $\Delta t \rightarrow 0$, one gets the stochastic Malthusian growth model. We denote $Y(t)=D^{m}(t)-$ $D^{f}(t)$, where $D^{m}(t)$ and $D^{f}(t)$ are the age specific mortality rates for males and females.

Consider a Stochastic differential equation (SDE)

$$
\begin{equation*}
\frac{d Y_{t}}{d t}=b\left(t, Y_{t}\right)+\sigma\left(t, Y_{t}\right) W_{t}, \quad b(t, y) \epsilon R \quad \sigma(t, y) \epsilon R \tag{1}
\end{equation*}
$$

Where $W_{t}$ is one-dimensional 'white noise'
By Itô (1951) calculus, the $Y_{t}$ satisfies the following Stochastic integral equation
$Y_{t}=Y_{0}+\int_{0}^{t} b\left(s, Y_{s}\right) d s+\int_{0}^{t} \sigma\left(t, Y_{s}\right) d s$
Or in differential form
$d Y_{t}=b\left(t, Y_{t}\right) d t+\sigma\left(t, Y_{t}\right) d B_{t}$
It is the $I t o ̂$ formula that is the key to the solution of many stochastic differential equations.
Various models describe the pattern of human mortality, the one such model that we consider is a stochastic Gompertz model also called Malthusian growth model.
$d Y(t)=R Y(t) d t+\sigma Y(t) d W(t), \quad Y(0)=Y_{0}>0$
Where $R$ is the arithmetic average growth rate, which is assumed constant over time and $W(t)$ is the white noise. The parameter $R$ was interpreted in the literature as an average growth rate. We have used Itô's calculus to obtain the solution. The solution is the geometric Brownian motion process.

The solution of equation (3) is obtained by Itô calculus (see also Øksendal 2003)
$Y(t)=y_{0} \exp \left\{\left(R-\frac{\sigma^{2}}{2}\right) t+\sigma W(t)\right\}$
or
$Y(t)=y_{0} \exp \{r t+\sigma W(t)\}$
Where $r=R-\frac{1}{2} \sigma^{2}$ is called the geometric average growth rate.
We conclude that

1) If $R>\frac{1}{2} \sigma^{2}$, then $Y(t) \rightarrow \infty$ as $t \rightarrow \infty$ a.s
2) If $R<\frac{1}{2} \sigma^{2}$, then $Y(t) \rightarrow 0$ as $t \rightarrow \infty$ a.s
3) If $R=\frac{1}{2} \sigma^{2}$, then $Y(t)$ will fluctuate between arbitrary large and small values as $t \rightarrow \infty$ a.s

To fit model (4), we made it age-specific by writing

$$
\begin{equation*}
Y_{x}(t)=y_{x 0} \exp \left\{\left(R_{x}-\frac{\sigma_{x}^{2}}{2}\right) t+\sigma_{x} W(t)\right\} \tag{6}
\end{equation*}
$$

Where $R_{x}=r_{x}+\frac{\sigma_{x}^{2}}{2}$ and hence, $r_{x}=R_{x}-\frac{\sigma_{x}^{2}}{2}$, which called as geometric average growth rate.
We apply the methodology used in (Lagarto and Braumann, 2014; Braumann, 2019; Talawar and Agadi, 2020) for parameters estimation. The value of $r_{x}$ of the rate of decay of $Y_{x}(t)$ would be negative since the tendency is for $Y_{x}(t)$ to decrease exponentially. The parameter $\sigma_{x}$ measures the intensity of the effect of environmental fluctuations (weather conditions, epidemic diseases, social conditions, etc.) on the rate of change. In this case of mortality rates are time equidistant observations with $\Delta=1$ year and $t_{k}=k \Delta$ years ( $k=0,1,2 \ldots n=60$ ), that is, the mortality rates are age-specific with single year. Denoting by $Y_{x, k}=Y_{x}\left(t_{k}\right)$ and by $L_{x, k}=\ln \left(Y_{x, k} / Y_{x,(k-1)}\right)$, the log-returns, one gets for model (6) the maximum likelihood estimates of the parameters:

$$
\hat{r}_{x}=\frac{1}{n \Delta} \sum_{k=1}^{n} L_{x, k} / \text { year and } \hat{\sigma}_{x}^{2}=\frac{1}{n \Delta} \sum_{k=1}^{n}\left(L_{x, k}-\hat{r}_{x} \Delta\right)^{2} / \text { year }
$$

Therefore the $95 \%$ confidence intervals of the parameters $r_{x}$ and $\sigma_{x}{ }^{2}$ are

$$
\hat{r}_{x} \pm 1.96 \sqrt{\hat{\sigma}_{x}^{2} / n \Delta} \text { and } \hat{\sigma}_{x}^{2} \pm \sqrt{2 \hat{\sigma}_{x}^{4} / n}
$$

## 3. APPLICATIONS OF THE MODEL

For the application of model we considered the mortality data of New Zealand from 1948-2015. In Figure 1 it shows that age-specific death rates of 60 -year-old females and males of the New Zealand population for each year of the period 1948-2015 (Source: https://www.mortality.org/). For convenience, we start counting time in 1948, so initial time, $t=0$ corresponds to 1948.

Figures 2-4, show the results of using model (6) for the age-specific mortality rates of 60 year-old New Zealand males and females. Notice the decline from $Y_{60}(0)=0.0138 /$ yearin 1948 to $Y_{60}(60)=0.00489 /$ year in 2008 and $Y_{60}(75)=0.00377 /$ year in 2023. For parameter estimation, we only use the data of the period 1948-2008, reserving the period 2009-2023 for prediction purposes.

Therefore the estimates of the parameters are

$$
\begin{aligned}
& \hat{r}_{60}=\frac{1}{n \Delta} \sum_{k=1}^{n} L_{60, k}=\frac{1}{n \Delta} \ln \frac{Y_{60, n}}{Y_{60}(0)} / \text { year } \\
& \hat{\sigma}_{60}{ }^{2}=\frac{1}{n \Delta} \sum_{k=1}^{n}\left(L_{60, k}-\hat{r}_{60} \Delta\right)^{2} / \text { year }
\end{aligned}
$$

The $95 \%$ confidence intervals of the parameters are
$\hat{r}_{60} \pm 1.96 \sqrt{\hat{\sigma}_{60}{ }^{2} / n \Delta}=(-1.0373 \pm 0.3371 /$ year $)$ and
$\hat{\sigma}_{60}{ }^{2} \pm \sqrt{2 \hat{\sigma}_{60}{ }^{4} / n}=(0.0291 \pm 0.003756) /$ year.
The parameter value $\hat{r}_{60}$ obtained is used in (6) to get the plots and putting $\sigma_{x}=0$, as if we have a deterministic environment, we obtain an adjusted curve (the thin solid line in Figures 1-7) that gives an idea of the trend. If one has the data up to 2008 (corresponding to $n=60$ and $t_{n}=2008-1948=60$ years) and wants to make predictions for the future period 2009-2023, one uses model (6) with the estimated parameter values (based only on data up to 2009), starting with the observed mortality rate of $2009, Y_{60}\left(t_{n}\right)=Y_{60}(60)=0.00489 /$ year. It is easy to work with $Z_{60}(t)=\ln Y_{60}(t)$, since it is Gaussian. For $\tau>0$ years, using the same technique as in Braumann (2019), one gets the point predictor
$\hat{Z}_{60}\left(t_{n}+\tau\right)=\ln 0.00489+\hat{r}_{60} \tau$
$\hat{Z}_{60}\left(t_{n}+\tau\right)=\exp \left(\hat{Z}_{60}\left(t_{n}+\tau\right)\right)$
and an approximate $95 \%$ confidence prediction interval for $\hat{Z}\left(t_{n}+\tau\right)$
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$\hat{Z}_{60}\left(t_{n}+\tau\right) \pm 1.96 \sqrt{\hat{\sigma}^{2} \tau\left(1+\tau / t_{n}\right.}$
From which extremes, by taking exponentials, one gets an approximate confidence prediction interval for $\hat{Y}_{60}\left(t_{n}+\tau\right)$. The similar procedure can be used to estimate $r_{x}$ and $\sigma_{x}{ }^{2}$ at each age $x$. The three lines in Figure 1-7 show the extremes of such a prediction interval (outer lines) and the point prediction (middle line). As we can see, the point predictions quite good, almost undistinguishable from the true mortality rates observed in 2009-2023.


Figure 1: Forecasted mortality rate of New Zealand male and female


Figure 2: Forecasted mortality rate of New Zealand female


Figure 3: Forecasted mortality rate of New Zealand male


Figure 4: Forecasted mortality rate of New Zealand male-female difference


Figure 5: Forecasted mortality rate of New Zealand Infant


Figure 6: Forecasted mortality rate of New Zealand at age 40


Figure 7: Forecasted mortality rate of New Zealand at age 80
Table 1: The parameter values for males and females of New Zealand mortality data.

|  | $Y(0) /$ year | $Y\left(t_{n}\right) /$ year | $Y(75) /$ year | $\hat{r} /$ year | $\hat{\sigma}^{2} /$ year |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $N Z_{60}^{f}$ | 0.0138 | 0.00489 | 0.00377 | -0.0173 | 0.0291 |
| $N Z_{60}^{m}$ | 0.02014 | 0.00792 | 0.00626 | -0.0156 | 0.01586 |
| $N Z_{\text {In }}^{T}$ | 0.02817 | 0.00463 | 0.00296 | -0.0301 | 0.0051 |
| $N Z_{40}^{T}$ | 0.00242 | 0.00098 | 0.00021 | -0.1051 | 0.0198 |
| $N Z_{80}^{T}$ | 0.10842 | 0.0499 | 0.04111 | -0.0129 | 0.0051 |
| $N Z_{60}^{D}$ | 0.00634 | 0.00303 | 0.00253 | -0.01231 | 0.1421 |

## 4. CONCLUSIONS

This stochastic mortality model gives good prediction intervals and the point prediction for each age. The point predictions of mortality rates at ages 40,60 and 80 are quite good, almost undistinguishable from the true mortality rates observed. The similar procedure can be used to estimate parameters at each age $x$. Once all the age-specific mortalities are obtained, the different columns of a life table can be constructed using their interrelationships.

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## CONFLICT OF INTEREST

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